# Probability, Random Variables, Distributions

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# **Building Blocks**

- Outcome: A specific result
  - One coin flip: Outcome is either *H* or *T*
  - Two coins flipped in sequence: HT is an outcome
- Sample space: The set of all possible outcomes, often denoted S
  - One coin flip:  $S_1 = \{H, T\}$
  - Two coins flipped in sequence:  $S_2 = \{HH, HT, TH, TT\}$
- Events: Subset of possible outcomes, a subset of S (can be S itself).
  - $\blacktriangleright A_1 = \{H\} \subseteq S_1; A_2 = \{HH, HT\} \subseteq S_2$
- Probability: Chance of an event within the sample space
  - ▶ A function that maps events to [0, 1]
  - Will give a more formal definition

## **Different Sizes of Sets**

Roughly, intuition tells us that

Probability =  $\frac{\text{Size of Event}}{\text{Size of Sample Space}} = \frac{\# \text{Outcomes}}{\# \text{All possible outcomes}}$ 

- Sample space can be of different sizes, leading to different treatments of probability
- Finite:
  - Ex. Number of seats in US House,  $S = \{0, 1, \dots, 435\}$
- Infinite but Countable:
  - Ex. Potential number of wars,  $S = \{0, 1, 2, 3, ...\}$
- Infinite and Uncountable:
  - Ex. Time duration of cabinets,  $S = [0, \infty)$
  - You cannot "count" the number of outcomes in this case

### **Set Operations**

Given two sets *A* and *B*, we can do the following operations (draw Venn diagrams): 1. **Union**: The set containing all of the elements in *A* **or** *B* 

$$A \cup B = \{ \omega : \omega \in A \text{ or } \omega \in B \}$$

2. Intersection: The set containing all of the elements in both A and B

 $A \cap B = \{ \omega : \omega \in A \text{ and } \omega \in B \}$ 

3. Complement: The set containing all of the elements not in A

$$A^c = \{ \omega \, : \, \omega \notin A \}$$

Other useful concepts:

- Set Difference:  $B \setminus A = \{ \omega : \omega \in B \text{ and } \omega \notin A \} = B \cap A^c$
- **Disjoint Sets**: A and B are disjoint if  $A \cap B = \emptyset$
- Indicator Function:  $\mathbb{1}(\omega \in A) = \begin{cases} 1 & \text{if } \omega \in A \\ 0 & \text{if } \omega \notin A \end{cases}$

# **Probability Function**

#### Definition

A **function**  $\mathbb{P}$  which assigns events to  $\mathbb{R}$  is called a **probability function** if it satisfies the following Axioms of Probability (Kolmogorov 1933)

1. For any event 
$$A$$
,  $\mathbb{P}[A] \ge 0$ 

**2.** 
$$\mathbb{P}[S] = 1$$

3. If  $A_1, A_2, \dots$  are mutually disjoint, then

(non-negative)

(sum up to 1, normalization)

(disjoint  $\Rightarrow$  sum)

$$\mathbb{P}\left[\bigcup_{i=1}^{\infty} A_i\right] = \sum_{i=1}^{\infty} \mathbb{P}[A_i]$$

Axiom 3 imposes that probabilities are additive on disjoint events

If 
$$A \cap B = \emptyset$$
, then  $\mathbb{P}(A \cup B) = \mathbb{P}(A) + \mathbb{P}(B)$ 

• Ex:  $S = \{H, T\}, \mathbb{P}[H] = 0.6, \mathbb{P}[T] = 0.6$  is not a valid probability function

# **Probability Properties**

For events *A* and *B*, given the three axioms, we can show the following properties: 1.  $\mathbb{P}[A^c] = 1 - \mathbb{P}[A]$ 

- 2.  $\mathbb{P}[\emptyset] = 0$
- 3.  $\mathbb{P}[A] \leq 1$
- 4. Monotone Probability Inequality: If  $A \subseteq B$ , then  $\mathbb{P}[A] \leq \mathbb{P}[B]$
- 5. Inclusion-Exclusion Principle:

$$\mathbb{P}[A \cup B] = \mathbb{P}[A] + \mathbb{P}[B] - \mathbb{P}[A \cap B]$$

6. Boole's Inequality:

$$\mathbb{P}[A \cup B] \le \mathbb{P}[A] + \mathbb{P}[B]$$

7. Bonferroni's Inequality:

 $\mathbb{P}[A \cap B] \ge \mathbb{P}[A] + \mathbb{P}[B] - 1$ 

# Joint Probability and Marginalization

- Joint Event  $(A \cap B)$  or simply (A, B): The event that both A and B occur
- Joint Probability: Probability that joint event (A, B) occurs, denoted

 $\mathbb{P}[A, B]$ 

• Example: Flip a coin twice,  $\mathbb{P}[H] = 0.6$ , A: 1st coin, B: 2nd coin

• **Marginalization**: We can recover the (marginal) probability  $\mathbb{P}[X]$  from joint probability  $\mathbb{P}[X, Y]$  by summing over every possible values of *Y*:

$$\mathbb{P}[X] = \sum_{y} \mathbb{P}[X, Y = y]$$

# **Conditional Probability**

• Conditional Probability: The conditional probability  $\mathbb{P}[A | B]$  is the probability of A given that B has occurred

$$\mathbb{P}[A | B] = \frac{\mathbb{P}[A, B]}{\mathbb{P}[B]}$$

- Allows for the inclusion of other information B into the calculation of probability of A
- Can think of B as the new sample space, and re-normalize all probabilities by  $\mathbb{P}[B]$
- Conditional probability is still a valid probability function (satisfies three axioms)
- This implies the Product Rule of probability: joint = conditional \* marginal

$$\mathbb{P}[A,B] = \mathbb{P}[A \mid B] \ \mathbb{P}[B]$$

or more generally (no particular order is needed),

$$\begin{split} \mathbb{P}[A, B, C, D, \ldots] &= \mathbb{P}[A] \ \mathbb{P}[B \mid A] \ \mathbb{P}[C \mid A, B] \ \mathbb{P}[D \mid A, B, C] \cdots \\ &= \mathbb{P}[D] \ \mathbb{P}[C \mid D] \ \mathbb{P}[B \mid C, D] \ \mathbb{P}[A \mid B, C, D] \cdots \end{split}$$

- **Independence**: The occurrence or nonoccurrence of either events *A* and *B* have no effect on the occurrence or nonoccurrence of the other; they are **unrelated**.
- The following are equivalent definitions for independence:
  - 1.  $\mathbb{P}[A | B] = \mathbb{P}[A]$
  - 2.  $\mathbb{P}[B|A] = \mathbb{P}[B]$
  - 3.  $\mathbb{P}[A, B] = \mathbb{P}[A] \mathbb{P}[B]$
- Conditioning on the event B does not modify the evaluation of probability of A
- Independent events provide no information to each other
- So conditional probability = unconditional probability

# **Conditional Independence**

- **Conditional Independence**: If *A* and *B* are independent once you know the occurrence of a **third event** *C*, then we say that *A* and *B* are conditionally independent **given** *C*.
- The following are equivalent definitions for conditional independence:
  - 1.  $\mathbb{P}[A | B, C] = \mathbb{P}[A | C]$
  - 2.  $\mathbb{P}[B|A,C] = \mathbb{P}[B|C]$
  - 3.  $\mathbb{P}[A, B | C] = \mathbb{P}[A | C] \mathbb{P}[B | C]$
- This is simply the definitions for independence but adding " $[\cdot | C]$ "
- This is a somewhat weaker condition than independence since we only need the above equality to hold on some subset of sample space involving event *C* 
  - But independence does not imply conditional independence, or vice versa
- This is one of the foundations of causal inference:  $\mathbb{P}[Y | D, X] = \mathbb{P}[Y | X]$ 
  - ▶ Given covariate X, treatment assignment D is independent of potential outcome Y

# **Probability: Example**

- A box contains two coins: a regular coin and one fake two-headed coin  $(\mathbb{P}[H] = 1)$ .
- I choose a coin at random  $\mathbb{P}[C] = p$  and toss it twice. Define the following events:
  - A = First coin toss results in an H
  - B = Second coin toss results in an H
  - C = Regular coin has been selected
- Find the following quantities:
  - $\blacktriangleright \mathbb{P}[A | C]$
  - $\blacktriangleright \mathbb{P}[B|C]$
  - $\blacktriangleright \mathbb{P}[A, B | C]$
  - $\mathbb{P}[A]$
  - ▶  $\mathbb{P}[B]$
  - $\blacktriangleright \mathbb{P}[A,B]$
- A and B are conditional independent given C, but A and B are not independent

## **Bayes Rule**

• From the **product rule** we know that

$$\mathbb{P}[A,B] = \mathbb{P}[B|A] \ \mathbb{P}[A] = \mathbb{P}[A|B] \ \mathbb{P}[B]$$

• From marginalization we also know that

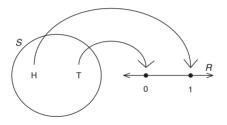
$$\mathbb{P}[B] = \mathbb{P}[A, B] + \mathbb{P}[A^c, B]$$
$$= \mathbb{P}[B | A] \mathbb{P}[A] + \mathbb{P}[B | A^c] \mathbb{P}[A^c]$$

• So we can express  $\mathbb{P}[A | B]$  as a function of  $\mathbb{P}[B | A]$  (and vise versa):

$$P[A | B] = \frac{P[A, B]}{P[B]} = \frac{P[B | A] P[A]}{P[B]}$$
$$= \frac{P[B | A] P[A]}{P[B | A] P[A] + P[B | A^c] P[A^c]}$$

#### **Random Variables**

• **Random Variable**: A random variable is a real-valued outcome; a function from the sample space S to real numbers  $\mathbb{R}$ 



• Example: Coin flip, we often use

$$X = \begin{cases} 1 & \text{if } H \\ 0 & \text{if } T \end{cases}$$

- X denote random variable
- X = x denote X has a particular realization x
- **Support** of *X*: The set that random variable *X* is defined, denoted  $\mathcal{X}$
- **Discrete** Random Variable: Sample space / support of X is finite or countable
- Continuous Random Variable: Sample space / support of X is uncountable

### **Distribution Function**

- We can then associate random variables with probability!
- **Cumulative Distribution Function (CDF)** of a random variable *X* is the probability that *X* is less than or equal to some value *x*:

 $F_X(x) = \mathbb{P}[X \le x]$ 

- A CDF F(x) must satisfy the following conditions:
  - 1. F(x) is non-decreasing in x (because we're including more outcomes)
  - 2.  $\lim_{x \to -\infty} F(x) = 0$  (probability of empty set)
  - 3.  $\lim_{x \to \infty} F(x) = 1$  (probability of whole sample space)
  - 4. F(x) is right-continuous (right limit must exist)

(technical)

- Continuous random variable ⇔ CDF is a continuous function
- Discrete random variable ⇔ CDF is a step function

#### **Density and Mass Functions**

• For continuous random variable, its Probability Density Function (PDF) is

$$f(x) = F'_X(x) = \frac{d}{dx}F_X(x)$$

- Fundamental Theorem of Calculus says  $\mathbb{P}[a \le X \le b] = \int_a^b f(x) dx = F(b) F(a)$
- For discrete random variable, its Probability Mass Function (PMF) is

$$f(x) = \mathbb{P}[X = x] \qquad \qquad [= F_X(x) - F_X(x - \varepsilon)]$$

- Since discrete CDF is a step function, it is not differentiable everywhere
- But we can still calculate that  $\mathbb{P}[a \le X \le b] = \sum_{x \in \{a, \dots, b\}} f(x) = \sum_{x \in \{a, \dots, b\}} \mathbb{P}[X = x]$
- Either way, PDF and PMF must satisfy the following conditions:
  - 1.  $f(x) \ge 0$  for all x (positivity)

2. 
$$\int_{-\infty}^{\infty} f(x) dx = 1 \text{ (PDF)} \quad \text{or} \quad \sum_{x} f(x) = 1 \text{ (PMF)} \quad (\text{sums to 1})$$

#### **Example: Bernoulli Random Variable**

•  $X \sim \text{Bernoulli}(p)$  if X has PMF

$$\mathbb{P}[X=1] = p$$
$$\mathbb{P}[X=0] = 1 - p$$

- The support of X is  $\{0, 1\}$
- Note that we can also write the PMF as

$$\mathbb{P}[X = x \mid p] = p^{x}(1-p)^{1-x}\mathbb{1}(x \in \{0, 1\})$$

• What is the CDF?

# **Example: Uniform Distribution**

•  $X \sim \text{Uniform}[0, 1]$  if X has PDF

$$f(x) = \begin{cases} 1 & 0 \le x \le 1\\ 0 & \text{otherwise} \end{cases}$$

- The support of X is [0, 1]
- We can also write the PDF as

$$f(x) = \mathbb{1}(x \in [0,1])$$

• What is the CDF?

# Expectation

- We often want to summarize characteristics of a distribution of a random variable
- What about we take the average of a random variable, weighted by probability?
- Expectation: The expected value of a random variable X is

$$\mathbb{E}[X] = \begin{cases} \int_{-\infty}^{\infty} xf(x) \, dx = \int_{-\infty}^{\infty} x \, dF(x) & \text{if } X \text{ continuous} \\ \sum_{x \in \mathcal{X}} xf(x) = \sum_{x \in \mathcal{X}} x \, \mathbb{P}[X = x] & \text{if } X \text{ discrete} \end{cases}$$

- Expectation tells us about the central tendency of a distribution
- More generally, we can find the expectation of any function of X

$$\mathbb{E}[g(X)] = \begin{cases} \int_{-\infty}^{\infty} g(x)f(x) \, dx = \int_{-\infty}^{\infty} g(x) \, dF(x) & \text{if } X \text{ continuous} \\ \sum_{x \in \mathcal{X}} g(x)f(x) = \sum_{x \in \mathcal{X}} g(x) \, \mathbb{P}[X = x] & \text{if } X \text{ discrete} \end{cases}$$

# Example

- Recall the Indicator Function:  $\mathbb{1}(\omega \in A) = \begin{cases} 1 & \text{if } \omega \in A \\ 0 & \text{if } \omega \notin A \end{cases}$
- Show that  $\mathbb{E}[\mathbb{1}(X \in A)] = \mathbb{P}[A]$

#### **Example: Bernoulli Random Variable**

•  $X \sim \text{Bernoulli}(p)$  if X has PMF

$$\mathbb{P}[X=1] = p$$
$$\mathbb{P}[X=0] = 1 - p$$

• What is  $\mathbb{E}[X]$ ?

• What is  $\mathbb{E}[X^2]$ ?

# **Example: Uniform Distribution**

•  $X \sim \text{Uniform}[0, 1]$  if X has PDF

$$f(x) = \begin{cases} 1 & 0 \le x \le 1\\ 0 & \text{otherwise} \end{cases}$$

• What is  $\mathbb{E}[X]$ ?

• What is  $\mathbb{E}[X^2]$ ?

# **Properties of Expected Values**

1. Expectation of a constant is a constant

$$\mathbb{E}[c] = c$$

2. Constants come out

$$\mathbb{E}[cg(Y)] = c \mathbb{E}[g(Y)]$$

3. Expectation is **linear**: For any random variables  $Y_1, \ldots, Y_n$  (either dependent or independent),

$$\mathbb{E}[g(Y_1) + \dots + g(Y_n)] = \mathbb{E}[g(Y_1)] + \dots + \mathbb{E}[g(Y_n)]$$

4. If X and Y are **independent**, then the product is easy

 $\mathbb{E}[XY] = \mathbb{E}[X] \mathbb{E}[Y]$ 

5. Expected Value of Expected Values:

$$\mathbb{E}[\mathbb{E}[Y]] = \mathbb{E}[Y]$$

(since  $\mathbb{E}[Y]$  is a constant)

#### **Variance and Covariance**

• How do we measure the distance of a random variable X from its mean?

 $X - \mathbb{E}[X]$ 

- But we would like the "distance" to stay positive...
  - Absolute value ~> Hard to deal with; not differentiable
  - Square ~> Smoother, differentiable
- Variance tells us about the spread of the distribution around the center

$$\operatorname{Var}[X] = \mathbb{E}\left[ (X - \mathbb{E}[X])^2 \right] = \mathbb{E}[X^2] - \mathbb{E}[X]^2$$

 Covariance measures the co-movement of two random variables around their own centers

$$\mathbb{C}\operatorname{ov}[X,Y] = \mathbb{E}\left[(X - \mathbb{E}[X])(Y - \mathbb{E}[Y])\right] = \mathbb{E}[XY] - \mathbb{E}[X]\mathbb{E}[Y]$$
  
Normalize:  $\operatorname{SD}[X] = \sqrt{\operatorname{Var}[X]}, \mathbb{C}\operatorname{orr}[X,Y] = \mathbb{C}\operatorname{ov}[X,Y]/\sqrt{\operatorname{Var}[X]\operatorname{Var}[Y]}$ 

#### **Properties of Variance and Covariance**

- 1. Var[c] = 0
- 2.  $\operatorname{War}[a+bX] = b^2 \operatorname{War}[X]$
- 3.  $\mathbb{C}ov[a+bX, c+dY] = bd \mathbb{C}ov[X, Y]$
- 4.  $\mathbb{C}ov[X + Z, Y + W] = \mathbb{C}ov[X, Y] + \mathbb{C}ov[X, W] + \mathbb{C}ov[Z, Y] + \mathbb{C}ov[Z, W]$
- 5.  $\operatorname{War}[X + Y] = \operatorname{War}[X] + \operatorname{War}[Y] + 2\operatorname{Cov}[X, Y]$

#### **Example: Bernoulli Random Variable**

•  $X \sim \text{Bernoulli}(p)$  if X has PMF

$$\mathbb{P}[X=1] = p$$
$$\mathbb{P}[X=0] = 1 - p$$

• What is Var[X]?

# **Example: Uniform Distribution**

•  $X \sim \text{Uniform}[0, 1]$  if X has PDF

$$f(x) = \begin{cases} 1 & 0 \le x \le 1\\ 0 & \text{otherwise} \end{cases}$$

• What is  $\operatorname{Var}[X]$ ?

#### **Binomial Distribution**

- Suppose we repeat the Bernoulli trial for *n* times
- Each trial *i* follows the same distribution

 $X_i \sim \text{Bernoulli}(p)$ , so  $\mathbb{P}[X_i = 1] = p$  for all i

• Each trial also independent of each other

$$\mathbb{P}[X_i = x_i, X_j = x_j] = \mathbb{P}[X_i = x_i] \mathbb{P}[X_j = x_j]$$

- We want to count the number of successes, denoted by  $Y = X_1 + \dots + X_n$
- **One specific way** to obtain *y* number of success is:

$$\mathbb{P}[X_1 = 1, X_2 = 0, X_3 = 0, ..., X_n = 0] = \mathbb{P}[X_1 = 1] \mathbb{P}[X_2 = 0] ... \mathbb{P}[X_n = 0]$$
  
=  $\underbrace{\mathbb{P}[X_1 = 1] \mathbb{P}[X_3 = 1] ...}_{y \text{ terms}} \times \underbrace{\mathbb{P}[X_2 = 0] ... \mathbb{P}[X_n = 0]}_{n-y \text{ terms}}$   
=  $\underbrace{pp \cdots p}_{y \text{ terms}} \times \underbrace{(1 - p) \cdots (1 - p)}_{n-y \text{ terms}} = p^y (1 - p)^{n-y}$ 

### **Binomial Distribution**

- This is just one instance of *y* number of successes
- There are  $\binom{n}{y} = \frac{n!}{y!(n-y)!}$  instances that we can get y successes
- So we have the PMF of the Binomial Distribution

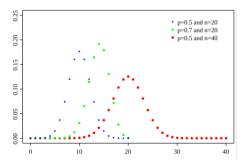
$$\mathbb{P}[Y = y | n, p] = \binom{n}{y} p^{y} (1 - p)^{n - y} \mathbb{1}(y \in \{0, 1, \dots, n\})$$

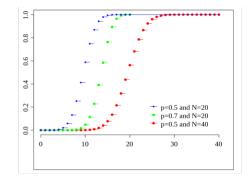
- The support of *Y* is  $\{0, 1, \dots, n\}$
- This example shows that:

If  $X_i \sim i.i.d.$  Bernoulli $(p) \Rightarrow X_1 + \dots + X_n = Y \sim \text{Binomial}(n, p)$ 

Note that the i.i.d. (independent and identically distributed) assumption is crucial in our derivation

### **Binomial PMF and CDF**





### **Binomial Distribution**

- $X_i \sim \text{ iid Bernoulli}(p), X_1 + \dots + X_n = Y \sim \text{Binomial}(n, p)$
- What is  $\mathbb{E}[Y]$ ?

• What is Var[Y]?

#### **Poisson Distribution**

#### Poisson distribution is often used to model rare event counts

- Counts of the number of events that occur during some unit of time
- The event would occur with a fixed "arrival rate"  $\lambda > 0$
- ex. Number of wars in a year (assuming "arrival rate" is fixed)
- $X \sim \text{Poisson}(\lambda)$  if X has PMF

$$\mathbb{P}[X = x \,|\, \lambda] = \frac{e^{-\lambda} \lambda^x}{x!} \mathbb{1}(x \in \{0, 1, 2, ...\})$$

• The support of *X* is {0, 1, 2, ...}

#### **Poisson Distribution**

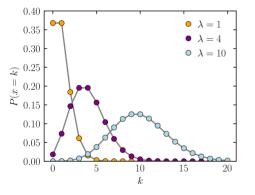
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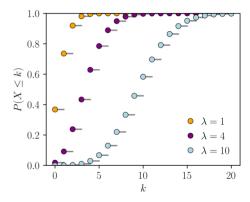
$$\mathbb{P}[X = x \,|\, \lambda] = \frac{e^{-\lambda} \lambda^x}{x!} \mathbb{1}(x \in \{0, 1, 2, ...\})$$

- Is this a legitimate PMF? (Does it follow the two conditions?)
  - Note the Taylor expansion of  $e^x$  around x = 0:

$$e^x = 1 + x + \frac{x^2}{2!} + \frac{x^3}{3!} + \cdots$$

#### **Poisson PMF and CDF**





#### **Poisson Distribution**

•  $X \sim \text{Poisson}(\lambda)$  if X has PMF

$$\mathbb{P}[X = x \,|\, \lambda] = \frac{e^{-\lambda} \lambda^x}{x!} \mathbb{1}(x \in \{0, 1, 2, ...\})$$

- What is  $\mathbb{E}[X]$ ?
  - Hint: Use the fact that PMF sums to 1:  $\sum_{x=0}^{\infty} \frac{e^{-\lambda} \lambda^x}{x!} = 1$

#### **Poisson Distribution**

•  $X \sim \text{Poisson}(\lambda)$  if X has PMF

$$P[X = x | \lambda] = \frac{e^{-\lambda} \lambda^x}{x!} \mathbb{1}(x \in \{0, 1, 2, ...\})$$

- Show that  $\mathbb{E}[X^2] = \lambda(\lambda + 1)$ 
  - Hint: Use the fact that PMF sums to 1:  $\sum_{x=0}^{\infty} \frac{e^{-\lambda} \lambda^x}{x!} = 1$

• Show that  $\operatorname{War}[X] = \lambda$